# NOTE ON A CASE-STUDY IN BOX-JENKINS SEASONAL FORECASTING OF TIME SERIES

BY

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TECHNICAL REPORT NO. 16
APRIL 1974

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FOR THE OFFICE OF NAVAL RESEARCH

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DEPARTMENT OF STATISTICS STANFORD UNIVERSITY STANFORD, CALIFORNIA Recently, the Journal of the Royal Statistical Society brought a lively discussion of the Box-Jenkins procedure for forecasting seasonal time series. This took place in two papers, one by Chatfield and Prothero (1973) and one by Box and Jenkins (1973). The discussion was centered around a particular case study of some sales figures.

Box and Jenkins refer to several other case-studies in their reply, including a paper on telephone installations and removals by Thompson and Tiao (1971). The original data in this analysis are not in the paper, but can be found in the preceding technical report from University of Wisconsin [Thompson and Tiao (1970)]. The present author had some time ago an opportunity to look at these data and the findings might be of some general interest in view of the recent discussion.

The data consist of the monthly totals of telephone installations and removals made by the Wisconsin Telephone Company from January 1951 through October 1966. The objective of the analysis is to obtain good forecasts for the two series. The authors consider the logarithms of the two series, respectively denoted by  $X_t$  and  $Y_t$ . Via the procedure described in the book by Box and Jenkins (1970) ARIMA models are obtained for  $Z_t = X_t - Y_t$  and  $Y_t$ . The parameters are estimated and forecasts made according to the estimated models. The forecasts of both series from November 1966 through October 1968 are compared to actual observations with quite good results. Following the standard notation the models for the  $Z_t$  and  $Y_t$  series were as follows:

$$(1 - B^{12}) (1 - \phi_1 B) Z_t = (1 - \theta_{12} B^{12}) (1 - \theta_1 B) a_t$$
, (1)

where  $a_{\pm}$  are uncorrelated random disturbances and

$$(1 - \phi_3 B^3) (1 - \phi_{12} B^{12}) Y_t = (1 - \delta_9 B^9 - \delta_{12} B^{12} - \delta_{13} B^{13}) b_t,$$
 (2)

The model (2) seems to be quite complicated, and (also suggested by Thompson and Tiao) calendar irregularities might be the reason. If one considers the number of installations or removals in any given month to be roughly proportional to the number of weekdays in that month, the effect of the calendar irregularities would be present in the  $Y_t$ -series but not in the  $Z_t$ -series as this involves only the <u>ratio</u> between the number of installations and removals. So as a simple method of forecasting the number of removals in any given month one might suggest that the average daily number of removals made in that month would increase by a fixed percentage over the corresponding number in the same month of the preceding year. That is, if we denote the number of weekdays in month t by  $d_t$ , one could let the forecast of  $Y_{T+12k+1}$ ,  $1 \leq j \leq 12$ , from  $Y_1, \ldots, Y_T$  be

$$\hat{Y}_{T+12k+j} = Y_{T+j-12} - \log d_{T+j-12} + \log d_{T+12k+j} + (k+1) \theta . \tag{3}$$

Averaging over  $(Y_t - \log d_t) = (Y_{t-12} - \log d_{t-12})$  suggests that  $\theta = 0.049$  would be a good value for  $\theta$ , corresponding to a 5% yearly increase.

The author tried to do exactly as in the paper by Thompson and Tiao apart from substituting the "naive" method (3) for the forecasts obtained from the model (2). That is  $Z_t$  was forecasted from the model (1) and  $Y_t$  by (3) and then  $X_t = Y_t + Z_t$  forecasted as

$$\hat{X}_{T+i} = \hat{Y}_{T+i} + \hat{Z}_{T+i} . \tag{4}$$

Tables 1 and 2 show the percentage error of forecasts for the total number of installations and removals, respectively, obtained by the two

methods for the 24 months from November 1966 through October 1968. These data are plotted on Figures 1 and 2.

Table 1

Percentage Error in Forecasts of the Number of Telephone Removals

Period	Nov. 1966		Jan. 1967	Feb.	Mar.	Apr.	May	June	July	Aug.	Sept.	Oct.
% error by original method	+6	+6	<b>-</b> 5	+3	+15	+12	-4	+1	+3	+7	+4	-7
% error by naive method	0	-1	0	+1	+6	+5	, 0	+1	+2	+8	0	<b>-</b> 2
						8 <sub>6</sub>						
Period	Nov.	Dec.	Jan. 1968	Feb.	Mar.	Apr.	May	June	July	Aug.	Sept.	Oct.
% error by original method	+8	+6	<b>-</b> 10	<b>-</b> 5	+14	+14	<b>-</b> 12	+7	<b>-</b> 12	+9	+7	0
% error by naive method	+8	-10	<b>-</b> 5	-7	+2	+17	-8	+8	-14	+5	+3	+9

Table 2

Percentage Error in Forecasts of the Number of Telephone Installations

Period	Nov. 1966		Jan. 1967	Feb.	Mar.	Apr.	May	June	July	Aug.	Sept.	Oct.
% error by original method	+9	+5	+2	+5	+9	+10	+2	<b>-</b> 2	+7	+7	+6	<b>-</b> 6
% error by naive method	+3	<b>-</b> 2	+7	+3	0	+3	+6	-2	+6	+8	+2	-1
						8	a .					
Period	Nov. 1967	Dec.	Jan. 1968	Feb.	Mar.	Apr.	May	June	July	Aug.	Sept.	Oct.
% error by original method	+8	+1	<b>-</b> 3	<b>-</b> 2	+14	+26	, <b>-</b> 10	+14	-11	0	+6	<b>-</b> 6
% error by naive method	+8	-15	+2	-4	+2	+29	, <b>-</b> 6	+15	-3	<b>−</b> jŧ	+2	+3

It is surprising but obvious that the "naive" method seems to give much better results for the forecasts of the removals and it does not do worse for the forecasts of the number of installations. None of the methods could forecast the strike in April 1968.

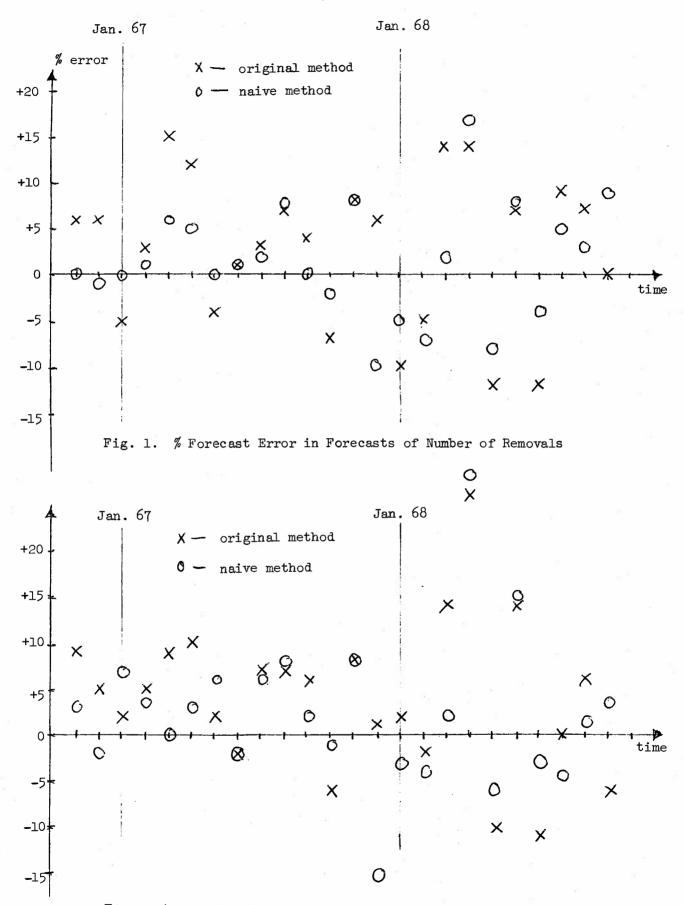


Fig. 2 % Forecast Error in Forecasts of Number of Installations

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It is shown that in a particular example a straight forward forecasting procedure gives better results than forecasts obtained from a fitted ARIMA model.